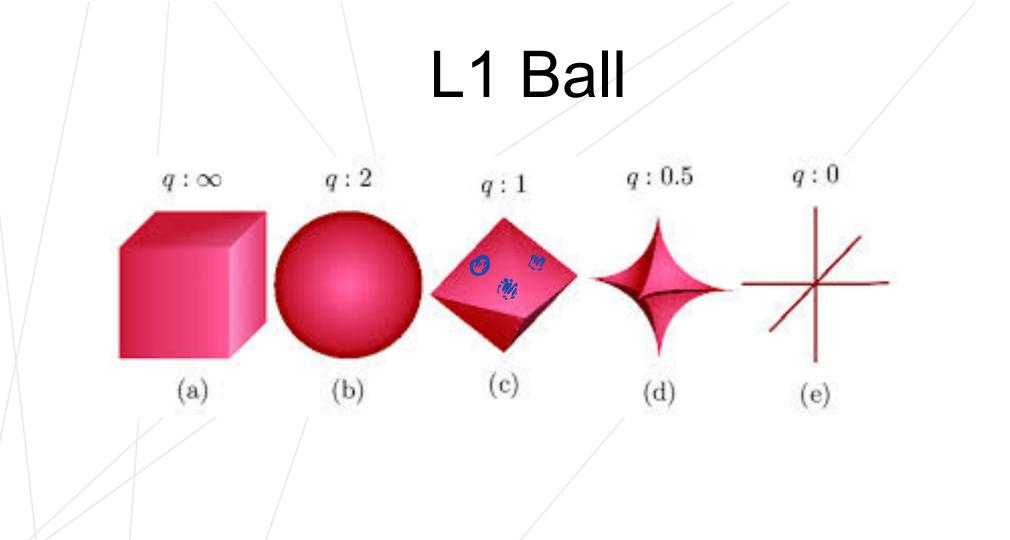
## Lecture 11 Localized Complexity

IEMS 402 Statistical Learning

# **Empirical Method of Maurey**



#### Volume Based Bound

Last lecture, we discussed the problem of getting a covering number N for  $L_1$  balls using  $L_2$  balls.

$$N(\epsilon, B_1^d, ||\cdot||_2) \tag{1}$$

Using a volume argument, we were able to establish the following result.

$$N(\epsilon, B_1^d, ||\cdot||_2) \le N(\epsilon, B_1^d, ||\cdot||_1)$$
 (2)

$$N(\epsilon, B_1^d, ||\cdot||_{\bullet}) \le (1 + \frac{2}{\epsilon})^d \tag{3}$$







## **Empirical Method of Maurey**

Theorem 1. When  $\epsilon > \frac{1}{\sqrt{d}}$ ,  $N \leq (2d+1)^{O(1/\epsilon^2)}$ 

As a result,  $\log N \lesssim \frac{1}{\epsilon^2} \log(d)$ .

*Proof.* Let's cover the following set:

$$B_1^{d,+} = \{ x \in \mathcal{R}^d \, | \, ||x||_1 \le 1 \text{ and } x_i \ge 0 \,\, \forall i \}$$

The above set means that  $\sum x_i \leq 1 \ \forall x_i \geq 0$ .

We can think about a probability distribution over  $\{e_1, \ldots, e_d, 0\}$ :

$$z=\sum_{i=1}^d x_i e_i + (1-\|x\|_1)\cdot 0$$
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## **Empirical Method of Maurey**

This implies the following probabilities.

$$\mathbb{P}[z = e_j] = x_j \,\forall j \in [d]$$
$$\mathbb{P}[z = 0] = 1 - ||x||_1$$

With these, we can get a mean of the probability distribution.

$$\mathbb{E}[z] = \sum \mathbb{P}[z = e_j] \cdot e_j + \mathbb{P}[z = 0] \cdot 0 = \sum x_j \cdot e_j = x$$

We will draw t samples  $z_1, \ldots, z_t$  from the distribution where each z is some  $e_i$ . After drawing the samples, we can take the average of the samples:

$$\bar{z} = \frac{1}{t} \sum_{i=1}^{t} z_i$$

We want to show that  $\mathbb{E}[\|\bar{z} - x\|_2^2] \leq \epsilon^2$ . If we can do this, then if we take all possible  $\bar{z}$ , we get an  $\epsilon$ -cover of the space using those  $\bar{z}$  since then all x we can choose will be within  $\epsilon$  of some point in the cover by what we argue above.

#### Empirical Method of Maurey vs Volumn

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The dependency on 2 is back

# Localized Complexity

### Example: Mean Estimation

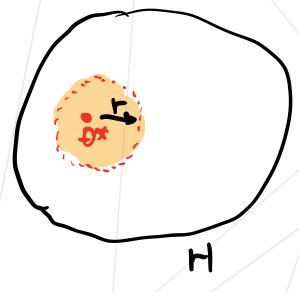
$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^{n} X_i = \underset{i=1}{\operatorname{argmin}} \frac{1}{n} (\theta - X_i)^2$$

$$X_i = \underset{i=1}{\operatorname{Min}} (\theta - X_i)^2$$

$$L(\theta) = \underbrace{E(\theta - X)^2}$$

$$L(\theta) =$$

### Idea:Localized Complexity



```
\phi(r) = \text{Rad} \left\{ \{ \theta \in H. \| \theta - \theta \| \leq r \} \right\}
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Now

#### Localize Leads to Fast Rate

## Non-parametric Least Square

To estimate the unknown regression function  $f^*$ , we consider the empirical risk minimizer (ERM), which is given by

$$\hat{f} = \underset{f \in \mathcal{F}}{\arg \min} \frac{1}{n} \sum_{i=1}^{n} (f(x_i) - y_i)^2.$$
 (2)

#### Method-

Since  $\hat{f}$  is optimal to the ERM problem (2) and  $f^* \in \mathcal{F}$  is feasible, we have **Proof** of Theorem 1:

$$\frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{f}(x_i))^2 \le \frac{1}{n} \sum_{i=1}^{n} (y_i - f^*(x_i))^2.$$
(3)

Also recall that

$$n = n$$

$$i=1$$

$$y_i = f^*(x_i) + \sigma w_i, \quad 1 \le i \le n.$$

We plug this expression into  $y_i$ 's in equation (3), open the squares and rearrange terms. Doing so gives the "basic inequality"

inequality"
$$\frac{1}{2}\|\hat{f} - f^*\|_n^2 \le \frac{\sigma}{n} \sum_{i=1}^n w_i(\hat{f}(x_i) - f^*(x_i)) \le \text{Red Corplex to }$$
(4)

Introducing the shorthand  $\Delta := \hat{f} - f^* \in \mathcal{F}^*$ , we rewrite the above basic inequality compactly as

$$\|\Delta\|^2 \leq C \cdot \|\Delta\|$$

$$= \frac{1}{2} \|\Delta\|_n^2 \leq \frac{\sigma}{n} \sum_{i=1}^n w_i \Delta(x_i) \cdot \Rightarrow \quad \text{ by the linear structs}$$

$$\Rightarrow \|\Delta\| \leq C \cdot \|\Delta\|_n^2 \leq \frac{\sigma}{n} \sum_{i=1}^n w_i \Delta(x_i) \cdot \Rightarrow \quad \text{ by the linear structs}$$

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#### We need star shape

localized  $G_n(S; T^*) := Rad(10-071 \le S, 0 \in H)$ 

**Lemma 1.** If  $\mathcal{F}^*$  is star-shaped, then the function  $\delta \mapsto \frac{G_n(\delta; \mathcal{F}^*)}{\delta}$  is non-increasing on  $(0, \infty)$ . Hence  $\delta^*$ **Proof** For any  $0 < \delta < t$ , we want to show that  $\frac{G_n(t,\mathcal{F}^*)}{t} \leq \frac{G_n(\delta;\mathcal{F}^*)}{\delta}$ .

Given  $h \in \mathcal{F}^*$  with  $||h||_n \leq t$ , define the rescaled function  $\tilde{h} = \frac{\delta}{t}h$ . We have  $\tilde{h} \in \mathcal{F}^*$  by definition with  $||h||_n \leq \delta$ . It is easy to see that

$$\frac{1}{n} \left( \frac{\delta}{t} \sum_{i=1}^{n} w_i h(x_i) \right) = \frac{1}{n} \sum_{i=1}^{n} w_i \tilde{h}(x_i).$$

Taking the supreme and expectation on both side over h, we obtain that

$$\frac{\delta}{t}\mathbb{E}\left[\sup_{h\in\mathcal{F}^*:\|h\|_n\leq t}\frac{1}{n}\sum_{i=1}^n w_ih(x_i)\right]\leq \mathbb{E}\left[\sup_{\tilde{h}\in\mathcal{F}^*:\|\tilde{h}\|_n\leq \delta}\frac{1}{n}\sum_{i=1}^n w_i\tilde{h}(x_i)\right].$$

This is equivalent to desired inequality

$$\frac{G_n(t, \mathcal{F}^*)}{t} \le \frac{G_n(\delta, \mathcal{F}^*)}{\delta}$$



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# Gn/S; F\*) = (82) -> the bu.

$$G_n(\delta; \mathcal{F}^*) = \delta$$

$$\Rightarrow F_{\bullet \bullet} \sup \frac{\sigma}{\sigma} \sum_{\sigma \in \mathcal{G}(x_{\bullet})} \sigma$$

$$\delta^* := \min_{\delta > 0} \left\{ \delta : \frac{G_n(\delta; \mathcal{F}^*)}{\delta} \le \frac{\delta}{2\sigma} \right\} \qquad \text{for } \sup_{\|g\|_n \le u} \frac{\sigma}{n} \sum_{i=1}^n \sigma_i g(x_i) \le u \delta^*$$

The If 
$$u \ge 8^*$$
, then  $E_{\sigma} \Omega P_{ilgM_n \le u} \overline{n} \Sigma \sigma_i g(x_i) \le u S^*$ .

Gin[ $u.F^*$ ) =  $u \cdot \frac{G_n(u.F^*)}{u} \le u \cdot \frac{G_n(8^*iF)}{8^*} = u \cdot \frac{8^*}{8^*}$ 

### Method 2: Peeling

**Lemma 1 (Peeling Technique)** If there is a function  $\phi:[0,\infty)\to [0,\infty)$  and  $r^*>0$  s.t.  $\forall r>\hat{r}^*$ , we have

- $\phi(4r) \leq 2\phi(r)$
- $R_n(G_r) \le \phi(r)$

Then we have for all  $r > \hat{r}^*$  we have

$$\mathbb{E}_{\sigma_i,z_i}\left[\frac{\frac{1}{n}\sum_{i=1}^n\sigma_ig(z_i)}{\mathbb{E}^{g+r}}\right] \leq \frac{4\phi(r)}{r} \quad \text{and} \quad \text{such that} \quad \phi(r^*) = r^*$$



